# Heavy-tailedness and diversification disasters: Implications for models in economics, finance and insurance 

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## Objectives and key results

- (Sub-)Optimality of diversification under heavy tails \& dependence
- (Non-)robustness of models in economics \& finance to heavy tails, heterogeneity \& dependence
- Implications for financial \& (re-)insurance markets: Diversification traps \& disasters
- M. Ibragimov, R. Ibragimov \& J. Walden, Heavy-tailedness and Robustness in Economics and Finance, Lecture Notes in Statistics, Springer, Forthcoming.
- R. Ibragimov \& A. Prokhorov, Topics in Majorization, Stochastic Openings and Dependence Modeling in Economics and Finance, World Scientific \& Imperial College Press, In preparation.


## Stylized Facts of Real-World Returns

Daily \% changes in the Dow Jones Industrial Average, Jan. 1980 - Sept. 2007


Dependence vs. margins in economic and financial problems

- Problems in finance, economics \& risk management:

Solution is affected by both

- Marginal distributions (Heavy-Tailedness, Skewness)
- Dependence (Positive or Negative, Asymmetry)
- Portfolio choice \& value at risk ( VaR )
- Marginal effects under independence: Heavy-Tailedness

Moderately HT vs. extremely $\mathrm{HT} \Longrightarrow$ Opposite solutions

- Different solutions: Positive vs. negative dependence
- Similar conclusions on (non-)robustness to heavy-tailedness: other models in economics, finance \& econometrics:
- Optimal bundling, firm growth theory, efficiency of statistical \& econometric estimators, time series models

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## Normal vs. Heavy-tailed Power Laws

Simulated normal and heavy-tailed series


## Heavy-tailed margins

- Many economic \& financial time series: power law tails: $P(|X|>x) \approx \frac{c}{x^{\alpha}}, \alpha>0$ : tail index
- Moments of order $p \geq \alpha$ : infinite; $E|X|^{p}<\infty$ iff $p<\alpha$
- $\alpha \leq 4 \Longrightarrow$ Infinite fourth moments: $E X^{4}=\infty$
- $\alpha \leq 2 \Longrightarrow$ Infinite variances: $E X^{2}=\infty$
- $\alpha \leq 1 \Longrightarrow$ Infinite first moments: $E|X|=\infty$
- Returns on many stocks \& stock indices: $\alpha \in(2,4)$
$\Rightarrow$ finite variance, infinite fourth moment


## A tale of two tails

Light vs. heavy tails


Figure: Tails of Cauchy distributions are heavier than those of normal distributions. Tails of Lévy distributions are heavier than those of Cauchy or normal distributions.

## A tale of two tails

Simulated data from Normal, Cauchy and Levy distributions, $\mathrm{n}=25$


Figure: Heavy-tailed distributions: more extreme observations

## Heavy-tailed margins

$P(|X|>x) \approx \frac{c}{x^{\alpha}}$

- Income: $\alpha \in[1.5,3] \Rightarrow$ infinite $E X^{4}$, possibly infinite variances
- Wealth: $\alpha \approx 1.5 \Rightarrow$ infinite variances!
- Returns from technological innovations, Operational risks: $\alpha<1 \Rightarrow$ infinite means $E|X|=\infty$ !
- Firm sizes, sizes of largest mutual funds, city sizes: $\alpha \approx 1$
- Economic losses from earthquakes: $\alpha \in[0.6,1.5]$
$\Rightarrow$ infinite variances, possibly infinite means
- Economic losses from hurricanes: $\alpha \approx 1.56 ; \alpha \approx 2.49$


## Stable distributions

- $X \sim S_{\alpha}(\sigma)$ : symmetric stable distribution, $\alpha \in(0,2]$

CF: $E\left(e^{i \times X}\right)=\exp \left\{-\sigma^{\alpha}|x|^{\alpha}\right\}$

- Normal $\mathcal{N}(0, \sigma): \alpha=2$
- Cauchy: $\alpha=1, f(x)=\frac{\sigma}{\pi\left(\sigma^{2}+x^{2}\right)}$
- Lévy: $\alpha=1 / 2$, support $[0, \infty), f(x)=\frac{\sigma}{\sqrt{2 \pi}} x^{-3 / 2} \exp \left(-\frac{1}{2 x}\right)$
- Power laws: $P(|X|>x) \approx \frac{c}{x^{\alpha}}, \alpha \in(0,2)$
- Moments $E|X|^{p}$ : finite iff $p<\alpha$
- Infinite variances for $\alpha<2$
- Portfolio formation: $\sum_{i=1}^{n} w_{i} X_{i}={ }_{d}\left(\sum_{i=1}^{n} w_{i}^{\alpha}\right)^{1 / \alpha} X_{1}$
- $\alpha=2$ (normal): $\frac{1}{\sqrt{n}}\left(X_{1}+\ldots+X_{n}\right)={ }_{d} X_{1}$


## Value at risk (VaR)

- VaR
- Risk $X$; positive values $=$ losses
- Loss probability $q$
- $\operatorname{Va} R_{q}(X)=z: P(X>z)=q$
- Risks $X_{1}, \ldots, X_{n}$
- $Z_{w}=\sum_{i=1}^{n} w_{i} X_{i}$ : return on portfolio with weights $w=\left(w_{1}, \ldots, w_{n}\right)$
- Problem of interest:

$$
\operatorname{Minimize} \operatorname{Va}_{a} R_{q}\left(Z_{w}\right)
$$

s.t. $w_{i} \geq 0, \sum_{i=1}^{n} w_{i}=1$

- When diversification $\Rightarrow$ decrease in portfolio riskiness (VaR)?


## Diversification \& risk

- Most diversified: $\underline{w}=(1 / n, 1 / n, \ldots, 1 / n) \Rightarrow Z_{\underline{w}}=\frac{1}{n} \sum_{i=1}^{n} X_{i}$
- Least diversified: $\bar{w}=(1,0, \ldots, 0) \Rightarrow Z_{\bar{w}}=X_{1}$
- $X_{1}, \ldots, X_{n} \sim \mathcal{N}(0, \sigma)(\alpha=2)$
- $Z_{\underline{w}}=\frac{1}{n} \sum_{i=1}^{n} X_{i}={ }_{d} \frac{1}{\sqrt{n}} X_{1}=\frac{1}{\sqrt{n}} Z_{\bar{w}}$
- $\operatorname{Va} R_{q}\left(Z_{\underline{w}}\right)=\frac{1}{\sqrt{n}} \operatorname{Va}_{q}\left(Z_{\bar{w}}\right)<\operatorname{Va}_{q}\left(Z_{\bar{w}}\right)$
- $\operatorname{Va} R_{q}\left(Z_{\underline{w}}\right): \searrow$ as $n \nearrow($ Diversification $\nearrow)$


## Diversification \& risk

- $X_{1}, \ldots, X_{n} \sim S_{1 / 2}(\sigma), \alpha=1 / 2$, Lévy distribution
- $Z_{\underline{w}}=\frac{1}{n} \sum_{i=1}^{n} X_{i}={ }_{d}\left[\sum_{i=1}^{n}\left(\frac{1}{n}\right)^{1 / 2}\right]^{2} X_{1}=n X_{1}=n Z_{\bar{w}}$
- $\operatorname{VaR}_{q}\left(Z_{\underline{\underline{w}}}\right)=n \operatorname{VaR}_{q}\left(Z_{\bar{w}}\right)>\operatorname{VaR}_{q}\left(Z_{\bar{w}}\right)$
- $\operatorname{VaR}_{q}\left(Z_{\underline{w}}\right): \nearrow$ as $n ~ \nearrow($ Diversification $\nearrow)$
- Heavy tails (margins) matter:
diversification $\Longrightarrow$ opposite effects on portfolio riskiness
- Skewness: typically priced


## Heavy-tailedness \& diversification

- Moderate heavy tails $\alpha>1$ : finite first moments

$$
\operatorname{Va}_{a}\left(Z_{\underline{w}}\right)<\operatorname{Va}_{q}\left(Z_{\bar{w}}\right) \quad \forall q>0
$$

Optimal to diversify for all loss probabilities $q$

- Extremely heavy tails $\alpha<1$ : infinite first moments

$$
\operatorname{Va}_{a}\left(Z_{\underline{w}}\right)<\operatorname{Va}_{q}\left(Z_{\bar{w}}\right) \quad \forall q>0
$$

Diversification: suboptimal for all loss probabilities $q$

- Similar conclusions: Many other models in economics \& finance
- Firm growth theory, optimal bundling, monotone consistency of sample mean, efficiency of linear estimators
- Robust to moderate heavy tails
- Properties: reversed under extremely heavy tails


## What happens for intermediate heavy-tails?

- $X_{1}, \ldots, X_{n}$ i.i.d. stable with $\alpha=1$ : Cauchy distribution
- Density $f(x)=\frac{\sigma}{\pi\left(\sigma^{2}+x^{2}\right)}$
- Heavy power law tails: $P(|X|>x) \approx \frac{c}{x}$
- Infinite first moment
- $Z_{w}=\sum_{i=1}^{n} w_{i} X_{i}={ }_{d} X_{1} \forall w=\left(w_{1}, \ldots, w_{n}\right): w_{i} \geq 0$,
- Diversification: no effect at all!


## Summary so far: Diversification for heavy-tailed and bounded distributions



Figure: $N=10$ risks/insurer; $M=7$ insurers

- D: Individual/non-diversification corners vs insurer and reinsurer equilibrium

1st example: full risk pooling with normally distributed risks

$1 \leq s \leq \frac{\text { Assume: }}{M(=5)}$ insurers
Results:
$N(=20)$ risks/insurer
$1 \leq j \leq N s$ total risks
i.i.d. normal $X_{i}$

CARA utility, Unlimited liability

$$
z_{j, s}=\left(\sum_{i=1}^{j} X_{i}\right) / s
$$

If no insurers pool, each still has $N$ risks

## 2nd example: Bernoulli-Lévy distribution with limited liability



Assume:
Limited liability: maximum loss ( $k=80$ ) $M=5$ insurers
$N(=20)$ max risks/insurer $u(x)=(x+k)^{3 / 4}$ $z_{j, s}=\left(\sum_{i=1}^{j} X_{i}\right) / s$

## Results:

If insurers can coordinate, they can reach $M N=100$ reinsurance equilibrium

But if not, each insurer reverts to the $N=0$ corner

## Implications for markets for catastrophic

## risks

- Equilibria in re-insurance markets for catastrophe risks (Ibragimov, Jaffee and Walden, RFS)
- A diversification equilibrium with full risk pooling for normally distributed (light-tailed) risks
- No risk pooling \& no insurance or reinsurance activity (market collapse) for extremely heavy-tailed cat risks
- Intermediate cases (heavy tails): both
- Diversification equilibria, in which insurers offer catastrophe coverage and reinsure their risks
- Non-diversification equilibria with no insurance or re-insurance
- A coordination problem must be solved to shift from the bad to the good equilibrium

Government regulations or well functioning capital markets

## Implications for markets for catastrophic

## risks

- Catastrophic risks have many features favorable to the provision of insurance
- Generally independent over risk types and geography
- Few issues of asymmetric information at the risk level
- So a complete failure of these markets is puzzling
- We have shown that market failures (non-diversification traps) may arise when risks are fat-tailed and there is limited liability
- Diversification may not be beneficial for the single insurer, although a full reinsurance equilibrium may exist.
- Government programs (or diversified equity owners) may allow the system to reach the full diversification outcome


## Diversification \& dependence

- Minimize $\operatorname{Va} R_{q}\left(w_{1} X_{1}+w_{2} X_{2}\right)$ s.t. $w_{1}, w_{2} \geq 0, w_{1}+w_{2}=1$
- Independence:
- Optimal portfolio: $\left(\tilde{w}_{1}, \tilde{w}_{2}\right)=\left(\frac{1}{2}, \frac{1}{2}\right)$ (diversified) if $\alpha>1$ (not extremely heavy-tailed, finite means)
- $\left(\tilde{w}_{1}, \tilde{w}_{2}\right)=(1,0)$ (not diversified, one risk) if $\alpha<1$ (extremely heavy-tailed, infinite means)


## Diversification \& dependence

- Extreme positive dependence: $X_{1}=X_{2}$ (a.s.) comonotonic risks
- $\operatorname{Va} R_{q}\left(w_{1} X_{1}+w_{2} X_{2}\right)=\operatorname{Va}_{q}\left(X_{1}\right) \forall w$
- Diversification: no effect at all (similar to Cauchy) regardless of heavy-tailedness
- Extreme negative dependence $X_{1}=-X_{2}$ (a.s.) countermonotonic risks
- $\operatorname{VaR} R_{q}\left(w_{1} X_{1}+w_{2} X_{2}\right)=\left(w_{1}-w_{2}\right) \operatorname{VaR}_{q}\left(X_{1}\right)$
- Optimal portfolio: $\underline{w}=(1 / 2,1 / 2)$ (most diversified regardless of heavy-tailedness
- Optimal portfolio choice: affected by both dependence \& properties of margins


## Copulas and dependence

- Main idea: separate effects of dependence from effects of margins
- What matters more in portfolio choice: heavy-tailedness \& skewness or (positive or negative) dependence?
- Copulas: functions that join together marginal cdf's to form multidimensional cdf


## Copulas and dependence

- Sklar's theorem
- Risks $X, Y$ :
- Joint cdf $H_{X Y}(x, y)=P(X \leq x, Y \leq y)$ : affected by dependence and by marginal cdf's $F_{X}(x)=P(X \leq x)$ and $G_{Y}(x)=P(Y \leq y)$
- $C_{X Y}(u, v)$ : copula of $X, Y$ :

$$
H_{X Y}(x, y)=\underbrace{C_{X Y}}_{\text {dependence }}(\underbrace{F_{X}(x), G_{Y}(y)}_{\text {marginals }})
$$

- $C_{X Y}$ : captures all dependence between risks $X$ and $Y$


## Copulas and dependence

Advantages:

- Exists for any risks (correlation: finiteness of second moments)
- Characterizes all dependence properties
- Flexibility in dependence modeling
- Asymmetric dependence: Crashes vs. booms
- Positive vs. negative dependence
- Independence: Nested as a particular case: Product copula, particular values of parameter(s)
- Extreme dependence: $X=Y$ or $X=-Y \Leftrightarrow$ extreme copulas; dependence in $C_{X Y}$ varies in between


## Copula structures

- Archimedean copulas

$$
C(u, v)=\phi^{-1}(\phi(u)+\phi(v))
$$

- Contagion: Non-zero tail dependence coeff.

$$
\begin{gathered}
\lambda_{L}=\lim _{u \rightarrow 0+} P\left[Y \leq F^{-1}(u) \mid X \leq F_{X}^{-1}(u)\right]=\lim _{u \rightarrow 0+} \frac{C(u, u)}{u} \\
\lambda_{U}=\lim _{u \rightarrow 1-} P\left[Y>F^{-1}(u) \mid X>F_{X}^{-1}(u)\right]=\lim _{u \rightarrow 1-} \frac{1-2 u+C(u, u)}{1-u}
\end{gathered}
$$

- Clayton \& Gumbel copulas


## Copula structures

- Eyraud-Farlie-Gumbel-Morgenstern (EFGM):

$$
C(u, v)=u v[1+\gamma(1-u)(1-v)]
$$

$\gamma \in[-1,1]$ : dependence parameter Tail independent: no contagion

- Heavy-tailed Pareto marginals:

$$
\begin{aligned}
& P(X>x)=\frac{1}{x^{\alpha}}, \quad x \geq 1 \\
& P(X>x)=\frac{1}{x^{\alpha}}, \quad x \geq 1
\end{aligned}
$$

- Power laws, tail index $\alpha$


## Diversification: Copulas \& heavy tails

Embrechts, Nešlehová \& Wüthrich (2009): Archimedean copulas

- Moderate heavy tails $\alpha>1$ : finite first moment

$$
\operatorname{Va}_{q}\left(\frac{X+Y}{2}\right)<\operatorname{Va}_{q}(X) \text { for sufficiently small } q
$$

Optimal to diversify for sufficiently small loss probabilities $q$

- Extremely heavy tails $\alpha<1$ : infinite first moments

$$
\operatorname{Va} R_{q}\left(\frac{X+Y}{2}\right)>\operatorname{Va}_{q}(X) \text { for sufficiently small } q
$$

Diversification: suboptimal for suff. small loss prob. $q$
Ibragimov \& Prokhorov (2013): Similar conclusions for EFGM

- Tail independent EFGM \& tail dependent Archimedean (Clayton, Gumbel): same boundary $\alpha=1$ as in the case of independence


## When dependence helps: Student- $t$ copulas

- Conclusions similar to independence: Models with common shocks

$$
X_{1}=Z Y_{1}, X_{2}=Z Y_{2}, \ldots, X_{n}=Z Y_{n}
$$

- Common shock $Z>0$ affecting all risks $X_{1}, \ldots, X_{n}$
- $Y_{1}, \ldots, Y_{n}$ : i.i.d. normal or heavy-tailed with tail index $\alpha$
$Z$ : heavy-tailed with tail index $\beta$
Then $X_{i}$ : heavy-tailed with tail index $\gamma=\min (\alpha, \beta)$
- Important particular case: (Dependent) Multivariate Student- $t$ $X_{1}, X_{2}, \ldots, X_{n}$ with $\alpha$ d.f. (tail index) $\Rightarrow \mathbf{O p t i m a l}$ to diversify for all loss probabilities $q$ regardless of tail index $\alpha$
- Tail dependent Student- $t$ copula and heavy-tailed margins with arbitrary tail index $\alpha$ : diversification pays off
- Contrast: Independent Student- $t X_{1}, X_{2}, \ldots, X_{n}$ with $\alpha$ d.f. (tail index): diversification optimal for $\alpha>1$; suboptimal for $\alpha<1$


## Diversification: Heavy-tailedness \& dependence matter

- Independence, Tail dependent models with common shocks (e.g., Student- $t$ distr. = Student- $t$ copula with Student- $t$ marginals):
- Diversification always pays off for all loss probabilities $q$
- Tail independent EFGM, possibly tail dependent Archimedean copulas (e.g., Clayton \& Gumbel):
- Dividing boundary $\alpha=1$ for sufficiently small loss probability $q$
- Numerical results on interplay of heavy-tailedness \& dependence (copula) assumptions and loss probability $q$ in diversification decisions:
- Deviations from threshold $\alpha=1$ for different copulas and loss probabilities $q$
- Theoretical results for general copulas $=$ ?
- (Non-)robustness of other models in economics \& finance


## Key results

- (Sub-)Optimality of diversification under heavy tails \& dependence
- (Non-)robustness of models in economics \& finance to heavy tails, heterogeneity \& dependence
- Implications for financial \& (re-)insurance markets: Diversification traps \& disasters
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## Characterizations of copulas \& dependence

- $V_{1}, \ldots, V_{n}$ : i.i.d. $\mathcal{U}([0,1])$
- C: $n$-copula iff $\exists \tilde{g}_{i 1}, \ldots,,_{c}$ s.t.

A1 (integrability):

$$
\int_{0}^{1} \ldots \int_{0}^{1}\left|\tilde{g}_{i_{1}, \ldots, i_{c}}\left(t_{i_{1}}, \ldots, t_{i_{c}}\right)\right| d t_{i_{1}} \ldots d t_{i_{c}}<\infty
$$

A2 (degeneracy):

$$
E_{V_{i_{k}}}\left[\tilde{g}_{i_{1}}, \ldots, i_{c}\left(V_{i_{1}}, \ldots, V_{i_{k-1}}, V_{i_{k}}, V_{i_{k+1}}, \ldots, V_{i_{c}}\right)\right]=0
$$

A3 (positive definiteness):

$$
\tilde{U}_{n}\left(V_{1}, \ldots, V_{n}\right) \equiv \sum_{c=2}^{n} \sum_{1 \leq i_{1}<\ldots<i_{c} \leq n} \tilde{g}_{i_{1}, \ldots, i_{c}}\left(V_{i_{1}}, \ldots, V_{i_{c}}\right) \geq-1
$$

- Representation for $C$ :

$$
C\left(u_{1}, \ldots, u_{n}\right)=\int_{0}^{u_{1}} \ldots \int_{0}^{u_{n}}\left(1+\tilde{U}_{n}\left(t_{1}, \ldots, t_{n}\right)\right) \prod_{i=1}^{n} d t_{i}
$$

- $\tilde{U}_{n}$ : sum of degenerate $U$-statistics

Device for constructing $n$-copulas and cdf's

- Bivariate Eyraud-Farlie-Gumbel-Morgenstern copulas \& cdf's:

$$
\begin{gathered}
C_{\theta}(u, v)=u v(1+\theta(1-u)(1-v)) \\
H_{\theta}(x, y)=F(x) G(y)(1+\theta(1-F(x))(1-G(y)) \\
n=2 ; \tilde{g}_{1,2}\left(t_{1}, t_{2}\right)=\theta\left(1-2 t_{1}\right)\left(1-2 t_{2}\right), \theta \in[-1,1]
\end{gathered}
$$

- Multivariate EFGM copulas \& cdf's:

$$
\begin{gathered}
C_{\theta}\left(u_{1}, u_{2}, \ldots, u_{n}\right)=\prod_{i=1}^{n} u_{i}\left(1+\theta \prod_{i=1}^{n}\left(1-u_{i}\right)\right) \\
\tilde{g}_{i_{1}, \ldots, i_{c}}\left(t_{i_{1}}, \ldots, t_{i_{c}}\right)=\theta_{i_{1}, \ldots, i_{c}}\left(1-2 t_{i_{1}}\right)\left(1-2 t_{i_{2}}\right) \ldots\left(1-2 t_{i_{c}}\right)
\end{gathered}
$$

- Generalized multivariate EFGM copulas (Johnson and Kotz, 1975, Cambanis, 1977)

$$
C\left(u_{1}, \ldots, u_{n}\right)=\prod_{k=1}^{n} u_{k}\left(1+\sum_{c=2}^{n} \sum_{1 \leq i_{1}<\ldots<i_{c} \leq n} \theta_{i_{1}, \ldots, i_{c}}\left(1-u_{i_{k}}\right)\right)
$$

$\tilde{g}_{i_{1}, \ldots, i_{c}}\left(t_{i_{1}}, \ldots, t_{i_{c}}\right)=0, c<n-1$
$\tilde{g}_{1,2, \ldots, n}\left(t_{1}, t_{2}, \ldots, t_{n}\right)=\theta\left(1-2 t_{1}\right)\left(1-2 t_{2}\right) \ldots\left(1-2 t_{n}\right)$

- Generalized EFGM copulas: complete characterization of joint cdf's of two-valued r.v.'s (Sharakhmetov \& Ibragimov, 2002)


## From dependence to independence through $U$-statistics

$\mathcal{G}_{n}$ : sums of $U$-statistics

$$
U_{n}\left(\xi_{1}, \ldots, \xi_{n}\right)=\sum_{c=2}^{n} \sum_{1 \leq i_{1}<\ldots<i_{c} \leq n} g_{i_{1}, \ldots, i_{c}}\left(\xi_{i_{1}}, \ldots, \xi_{i_{c}}\right)
$$

$g_{i 1}, \ldots, i_{c}:$ satisfy A1-A3

- Arbitrarily dependent r.v.'s:
sum of $U$-statistics in independent r.v.'s
with canonical kernels
- Reduction of problems for dependence to well-studied objects
- Transfer of results for $U$-statistics under independence


## From dependence to independence through

 $U$-statistics- $X_{1}, \ldots, X_{n}: 1$-cdf's $F_{k}\left(x_{k}\right)$
- $\xi_{1}, \ldots, \xi_{n}$ : independent copies (1-cdf's $\left.F_{k}\left(x_{k}\right)\right)$
$\exists U_{n} \in \mathcal{G}_{n}$ s.t. $\forall f: \mathbf{R}^{n} \rightarrow \mathbf{R}$

$$
E f\left(X_{1}, \ldots, X_{n}\right)=E f\left(\xi_{1}, \ldots, \xi_{n}\right)\left(1+U_{n}\left(\xi_{1}, \ldots, \xi_{n}\right)\right)
$$

- Representation for c.f.'s:

$$
\begin{gathered}
\operatorname{Eexp}\left(i \sum_{k=1}^{n} t_{k} x_{k}\right)=\operatorname{Eexp}\left(i \sum_{k=1}^{n} t_{k} \xi_{k}\right)+ \\
\operatorname{Eexp}\left(i \sum_{k=1}^{n} t_{k} \xi_{k}\right) U_{n}\left(\xi_{1}, \ldots, \xi_{n}\right)
\end{gathered}
$$

$\ddagger$ CLT for bivariate r.v.'s

## Characterizations of dependence

- Canonical $g^{\prime}$ s: complete characterizations of dependence properties
- $X_{1}, \ldots, X_{n}$ : $r$-independent if $\forall r$ jointly independent $\Leftrightarrow$ $g_{i_{1}, \ldots, i_{c}}\left(V_{i_{1}}, \ldots, V_{i_{c}}\right)=0$ (a.s.) $1 \leq i_{1}<\ldots<i_{c} \leq n, c=2, \ldots, r$
$g_{i_{1}, \ldots, i_{r+1}}\left(u_{i_{1}}, \ldots, u_{i_{r+1}}\right)=$
$\frac{\alpha_{1} \ldots \alpha_{n}}{\alpha_{i_{1}} \ldots \alpha_{i r+1}}\left((k+1) u_{i_{1}}^{k}-(k+2) u_{i_{1}}^{k+1}\right) \times \ldots \times\left((k+1) u_{i_{c}}^{k}-(k+2) u_{i_{c}}^{k+1}\right)$

$$
\begin{gathered}
C\left(u_{1}, \ldots, u_{n}\right)=\prod_{i=1}^{n} u_{i}\left(1+\sum_{1 \leq i_{1}<\ldots<i_{r+1} \leq n} \frac{\alpha_{1} \ldots \alpha_{n}}{\alpha_{1} \ldots \alpha_{i_{r}+1}} \times\right. \\
\left.\left(u_{i_{1}}^{k}-u_{i_{1}}^{k+1}\right) \times \ldots \times\left(u_{i_{r+1}}^{k}-u_{i_{r+1}}^{k+1}\right)\right)
\end{gathered}
$$

Extensions of Wang (1990) ( $k=0$ )

## Copulas and Markov processes

- Darsow, Nguyen and Olsen, 1992: copulas and first-order Markovness
- $A, B:[0,1]^{2} \rightarrow[0,1]:$

$$
(A * B)(x, y)=\int_{0}^{1} \frac{\partial A(x, t)}{\partial t} \cdot \frac{\partial B(t, y)}{\partial t} d t
$$

- $A:[0,1]^{m} \rightarrow[0,1], B:[0,1]^{n} \rightarrow[0,1]: \star-$ product

$$
\begin{gathered}
A \star B\left(x_{1}, \ldots, x_{m+n-1}\right)= \\
\int_{0}^{x_{m}} \frac{\partial A\left(x_{1}, \ldots, x_{m-1}, \xi\right)}{\partial \xi} \cdot \frac{\partial B\left(\xi, x_{m+1}, \ldots, x_{m+n-1}\right)}{\partial \xi} d \xi
\end{gathered}
$$

## Copulas and Markov processes

- Transition probabilities
$P(s, x, t, A)=P\left(X_{t} \in A \mid X_{s}=x\right)$ satisfy CKE's
iff $C_{s t}=C_{s u} * C_{u t} \forall s<u<t$
- $X_{t}$ : first-order Markov iff

$$
C_{t_{1}, \ldots, t_{n}}=C_{t_{1} t_{2}} \star C_{t_{2} t_{3}} \star \ldots \star C_{t_{n-1} t_{n}}
$$

## New results: Higher-order Markovness and copulas

- $\left\{X_{t}\right\}_{t \in T}: k$-order Markov $\Leftrightarrow$

$$
\begin{gathered}
P\left(X_{t}<X_{t} \mid X_{t_{1}}, \ldots, X_{t_{n-k}}, X_{t_{n-k+1}}, \ldots, X_{t_{n}}\right)= \\
P\left(X_{t}<x_{t} \mid X_{t_{n-k+1}}, \ldots, X_{t_{n}}\right)
\end{gathered}
$$

- Complete characterization in terms of ( $k+1$ )-copulas
- $C_{t_{1}, \ldots, t_{k}}$ : copulas of $X_{t_{1}}, \ldots, X_{t_{k}}$
- $\left\{X_{t}\right\}_{t \in T}: k$-order Markov iff $\forall t_{1}<\ldots<t_{n}, n \geq k+1$

$$
C_{t_{1}, \ldots, t_{n}}=C_{t_{1}, \ldots, t_{k+1}} \star^{k} C_{t_{2}, \ldots, t_{k+2}} \star^{k} \ldots \star^{k} C_{t_{n-k}, \ldots, t_{n}}
$$

## Stationary case

- $X_{t}$ : stationary $k$-order Markov iff

$$
\begin{gathered}
C_{1, \ldots, n}\left(u_{1}, \ldots, u_{n}\right)=C \star^{k} C \star^{k} \ldots \star^{k} C\left(u_{1}, \ldots, u_{n}\right) \\
=C^{n-k+1}\left(u_{1}, \ldots, u_{n}\right) \forall n \geq k+1
\end{gathered}
$$

C: $(k+1)-$ copula s.t.

$$
C_{i_{1}+h, \ldots, i_{i}+h}=C_{i_{1}, \ldots, i_{i}}, \quad 1 \leq j_{1}<\ldots<j_{1} \leq k+1
$$

- $C^{s}: s$-fold product $\star^{k}$ of $C$


## Advantages of copula-based approach

- Modeling higher order Markov processes
alternative to transition matrices
$\ddagger$ Instead of initial distribution \& transition probabilities:

Prescribe marginals \& $(k+1)$-copulas

Generate copulas of higher order \& finite-dimensional cdf's
$\ddagger$ Advantage: separation of properties of marginals (fat-tailedness) \& dependence properties (conditional symmetry, $m$-dependence, $r$-independence, mixing)

## Advantages of copula-based approach

- Inversion method:

New $k$-Markov with dependence similar to a given Markov process Different marginals
$\ddagger X_{t}$ : stationary $k$-Markov
$(k+1)-\operatorname{cdf} \tilde{F}\left(x_{1}, \ldots, x_{k+1}\right), 1-\operatorname{cdf} F$
$\Rightarrow(k+1)$-copula:

$$
C\left(u_{1}, \ldots, u_{k+1}\right)=\tilde{F}\left(F^{-1}\left(u_{1}\right), \ldots, F^{-1}\left(u_{k+1}\right)\right)
$$

$\dagger$ Another 1-cdf G:
Stationary $k$-Markov, same dependence as $\left\{X_{t}\right\}$, different 1-marginal $G$ :
$(k+1)$-copula:

$$
C\left(u_{1}, \ldots, u_{k+1}\right)=\tilde{F}\left(G^{-1}\left(u_{1}\right), \ldots, G^{-1}\left(u_{k+1}\right)\right)
$$

Representation $\Rightarrow$ Higher-order copulas \& cdf's
$\left\{X_{t}\right\}$ : stationary $C$-based $k$-Markov chain

## Advantages of copula-based approach

- C: all dependence properties of the time series
$\ddagger k$-independence, $m$-dependence, martingaleness, symmetry
$\ddagger$ On-going project with Johan Walden: characterizations of time-irreversibility; focus on $C_{t_{1}, \ldots, t_{k}}=C_{t_{k}, \ldots, t_{1}}$
$\ddagger$ Applications: forward-looking vs. backward-looking market participants ("fundamentalists" vs. noise traders or "chartists")
$\ddagger$ "Compass rose" for $P_{t-1}$ and $P_{t}$ : symmetry in copulas


## Combining higher-order Markovness with other dependence properties

- A number of studies in dependence modeling: Higher-order Markovness + $m$-dependence \& $r$-independence

Lévy (1949): 2nd order Markovness + pairwise independence

Rosenblatt \& Slepian (1962): $N$-order $N$-independent stationary Markov

- Impossibility/reduction :
$N$-order Markov $+N$-independence + two-valued $\Leftrightarrow$ joint independence
$\ddagger$ Testing sensitivity to WD in DGP Rosenblatt \& Slepian (1962)


## Combining Markovness with other dependencies

$\ddagger$ Examples:

Not 1-order Markovian

But 1-st order transition probabilities
$P(s, x, t, A)=P\left(X_{t} \in A \mid X_{s}=x\right)$ satisfy C-K SE

$$
P(s, x, t, A)=\int_{-\infty}^{\infty} P(u, \xi, t, A) P(s, x, u, d \xi)
$$

(other examples: Feller, 1959, Rosenblatt, 1960)

## Combining Markovness with other dependencies

$\ddagger$ 1-dependent Markov: Aaronson, Gilat and Keane (1992)

Burton, Goulet and Meester (1993), Matúš (1996)
$\ddagger$ Matúš (1998): m-dependent
discrete-space Markov
$\ddagger$ Impossibility/Reduction:
$\nexists$ stationary $m$-dependent Markov if
$\operatorname{card}(\Omega)<m+2$

## Markovness of higher-order and $k$-independence

- Characterization of stationary
$k$-independent $k$-Markov processes
- $\left\{X_{t}\right\}: C$-based $k$-independent stationary
$k$-Markov iff

$$
\frac{\partial^{k+1} C\left(u_{1}, \ldots, u_{k+1}\right)}{\partial u_{1} \ldots \partial u_{k+1}}=1+g\left(u_{1}, \ldots, u_{k+1}\right)
$$

$g:[0,1]^{k+1} \rightarrow[0,1]:$ canonical $g$-function
(Integrability + more degeneracy + positive definiteness)

## Markovness of higher-order and $k$-independence

$$
\begin{gathered}
\int_{0}^{1} \ldots \int_{0}^{1}\left|g\left(u_{1}, \ldots, u_{k+1}\right)\right| d u_{1} \ldots d u_{k+1}<\infty \\
\int_{0}^{1} \ldots \int_{0}^{1} g\left(u_{1}, \ldots, u_{k+1}\right) g\left(u_{2}, \ldots, u_{k+2}\right) \ldots g\left(u_{s}, \ldots, u_{k+s}\right) d u_{i_{1}} \ldots d u_{i_{s}}=0 \\
\forall s \leq u_{i_{1}}<\ldots<u_{i_{s}} \leq k+1, s=1,2, \ldots,\left[\frac{k+1}{2}\right] \\
g\left(u_{1}, \ldots, u_{k+1}\right) \geq-1
\end{gathered}
$$

- Integration: w.r. to all $s$ among $u_{s}, u_{s+1}, \ldots, u_{k+1}$ common to all $g$-functions $g\left(u_{1}, \ldots, u_{k+1}\right), g\left(u_{2}, \ldots, u_{k+2}\right), \ldots, g\left(u_{s}, \ldots, u_{k+s}\right)$
$k$-marginals: product copulas, independence
$k$-independence: satisfied


## Markovness of higher-order and $m$-independence

- $\left\{X_{t}\right\}: C$-based $m$-dependent 1 -Markov iff

$$
\frac{\partial^{2} C\left(u_{1}, u_{2}\right)}{\partial u_{1} \partial u_{2}}=1+g\left(u_{1}, u_{2}\right)
$$

$g:[0,1]^{2} \rightarrow[0,1]:$ canonical $g$-function:

$$
\begin{gathered}
\int_{0}^{1} \int_{0}^{1}\left|g\left(u_{1}, u_{2}\right)\right| d u_{1} d u_{2}<\infty \\
\int_{0}^{1} g\left(u_{1}, u_{2}\right) d u_{i}=0, \quad g\left(u_{1}, u_{2}\right) \geq-1 \\
\int_{0}^{1} g\left(u_{1}, u_{2}\right) g\left(u_{2}, u_{3}\right) \ldots g\left(u_{m}, u_{m+1}\right) d u_{2} d u_{3} \ldots d u_{m}=0
\end{gathered}
$$

$\ddagger$ Integration: w.r. to $u_{2}, u_{3}, \ldots, u_{m}$ more than once among $g\left(u_{1}, u_{2}\right), g\left(u_{2}, u_{3}\right)$,
$\ldots, g\left(u_{m}, u_{m+1}\right)$
$X_{1}, X_{m+1}$ : independent; Process: $m$-dependent

## New examples via existing constructions

- Higher-order Markovness + martingaleness
- Inversion method + existing examples $\Rightarrow$
$k$-independent, $m$-dependent Markov processes
different marginals


## Reduction \& impossibility for $k$-order Markov

 processes- $\left\{X_{t}\right\}$ : $C$-based $k$-independent stationary $k$-Markov
$\ddagger \frac{\partial^{k+1} c\left(u_{1}, \ldots, u_{k+1}\right)}{\partial u_{1} \ldots \partial u_{k+1}}=1+g\left(u_{1}, \ldots, u_{k+1}\right)$
$\ddagger g$ : product form (EFGM-type):
$g\left(u_{1}, u_{2}, \ldots, u_{k+1}\right)=\alpha f\left(u_{1}\right) f\left(u_{2}\right) \ldots f\left(u_{k+1}\right)$
$\Leftrightarrow\left\{X_{t}\right\}$ : jointly independent


## Examples: EFGM and power copulas

- $(k+1)$-EFGM copulas:

$$
\begin{gathered}
C\left(u_{1}, u_{2}, \ldots, u_{k+1}\right)=\prod_{i=1}^{k+1} u_{i}\left(1+\alpha\left(1-u_{1}\right)\left(1-u_{2}\right) \ldots\left(1-u_{k+1}\right)\right) \\
g\left(u_{1}, u_{2}, \ldots, u_{k+1}\right)=\alpha\left(1-2 u_{1}\right)\left(1-2 u_{2}\right) \ldots\left(1-2 u_{k+1}\right)
\end{gathered}
$$

- $(k+1)$-power copulas

$$
C\left(u_{1}, u_{2}, \ldots, u_{k+1}\right)=\prod_{i=1}^{k+1} u_{i}\left(1+\alpha\left(u_{1}^{\prime}-u_{1}^{I+1}\right)\left(u_{2}^{\prime}-u_{2}^{I+1}\right) \ldots\left(u_{k+1}^{\prime}-u_{k+1}^{I+1}\right)\right)
$$

$I \geq 0(E F G M: I=0)$

## Impossibility/reduction for $m$-dependence

- $\left\{X_{t}\right\}: C$-based $m$-dependent Markov
$\ddagger \frac{\partial^{2} C\left(u_{1}, u_{2}\right)}{\partial u_{1} \partial u_{2}}=1+\alpha f\left(u_{1}\right) f\left(u_{2}\right)$
(separable product form)
$\Leftrightarrow X_{t}$ : jointly independent
- Representations $\Rightarrow$

$$
\begin{gathered}
\int_{0}^{1} \ldots \int_{0}^{1} \alpha^{m} f\left(u_{1}\right) f^{2}\left(u_{2}\right) \ldots f^{2}\left(u_{m}\right) f\left(u_{m+1}\right) d u_{2} \ldots d u_{m}=0 \\
\alpha^{m} f\left(u_{1}\right) f\left(u_{m+1}\right)\left[\int_{0}^{1} f^{2}\left(u_{2}\right) d u_{2}\right]^{m-1}=0
\end{gathered}
$$

$\Rightarrow f=0 \Leftrightarrow$ Independence

## Examples, new and old

$\ddagger$ EFGM copulas, $k=1$ :

$$
\begin{gathered}
C\left(u_{1}, u_{2}\right)=u_{1} u_{2}\left(1+\alpha\left(1-u_{1}\right)\left(1-u_{2}\right)\right) \\
g\left(u_{1}, u_{2}\right)=\alpha\left(1-2 u_{1}\right)\left(1-2 u_{2}\right)
\end{gathered}
$$

- Limitations of EFGM copulas,
separable copulas:
Complement \& generalize existing results


## Examples, new and old

$\ddagger$ Cambanis (1991): common dependencies
cannot be exhibited by multivariate EFGM

$$
\begin{gathered}
C_{j_{1}, \ldots, j_{n}}\left(u_{j_{1}}, \ldots, u_{j_{n}}\right)= \\
\prod_{s=1}^{n} u_{j_{k}}\left(1+\sum_{1 \leq l<m \leq n} \alpha_{l m}\left(1-u_{j_{l}}\right)\left(1-u_{j_{m}}\right)\right)
\end{gathered}
$$

$\ddagger$ Rosenblatt \& Slepian (1962): non-existence of bivariate $N$-independent $N$-Markov

Sharakhmetov \& Ibragimov (2002):

## EFGM copulas for two-valued r.v.'s

$\ddagger$ Technical difficulties in modeling

## Solution: New flexible copula classes

- Copula-based TS with flexible dependencies
$\ddagger$ Copulas based on Fourier polynomials
- $k$-independent $k$-Markov: Conditions satisfied for

$$
g\left(u_{1}, \ldots, u_{k+1}\right)=\sum_{j=1}^{N}\left[\alpha_{j} \sin \left(2 \pi \sum_{i=1}^{k+1} \beta_{i}^{j} u_{i}\right)+\gamma_{j} \cos \left(2 \pi \sum_{i=1}^{k+1} \beta_{i}^{j} u_{i}\right)\right]
$$

$\ddagger \alpha_{j}, \gamma_{j} \in \mathbf{R}, \beta_{i}^{j} \in \mathbf{Z}, i=1, \ldots, k+1, j=1, \ldots, N:$
$\dagger \beta_{1}^{j_{1}}+\sum_{l=2}^{s} \epsilon_{l-1} \beta_{l}^{j_{l}} \neq 0$
$\epsilon_{1}, \ldots, \epsilon_{s-1} \in\{-1,1\}, s=2, \ldots, k+1$
$\dagger 1+\sum_{j=1}^{N}\left[\alpha_{j} \epsilon_{j}+\gamma_{j} \epsilon_{j+N}\right] \geq 0, \epsilon_{1}, \ldots, \epsilon_{2 N} \in\{-1,1\}$

## Fourier copulas

$$
C\left(u_{1}, \ldots, u_{k+1}\right)=\int_{0}^{u_{1}} \ldots \int_{0}^{u_{k+1}}\left(1+g\left(u_{1}, \ldots, u_{k+1}\right)\right) d u_{1} \ldots d u_{k+1}
$$

$(k+1)$-Fourier copulas

## Fourier copulas

- 1-dependent 1-Markov:

Conditions satisfied for Fourier copulas

$$
\begin{gathered}
C\left(u_{1}, u_{2}\right)=\int_{0}^{u_{1}} \int_{0}^{u_{2}}\left(1+g\left(u_{1}, u_{2}\right)\right) d u_{1} d u_{2} \\
g\left(u_{1}, u_{2}\right)=\sum_{j=1}^{N}\left[\alpha_{j} \sin \left(2 \pi\left(\beta_{1}^{j} u_{1}+\beta_{2}^{j} u_{2}\right)\right)+\gamma_{j} \cos \left(2 \pi\left(\beta_{1}^{j} u_{1}+\beta_{2}^{j} u_{2}\right)\right)\right] \\
\ddagger \alpha_{j}, \gamma_{j} \in \mathbf{R}, \beta_{1}^{j}, \beta_{2}^{j} \in \mathbf{Z}: \quad \beta_{1}^{j_{1}}+\beta_{2}^{j_{2}} \neq 0 \\
\beta_{1}^{j_{1}}-\beta_{2}^{j_{2}} \neq 0 \\
1+\sum_{j=1}^{N}\left[\alpha_{j} \epsilon_{j}+\gamma_{j} \epsilon_{j+N}\right] \geq 0 \\
\forall \epsilon_{1}, \ldots, \epsilon_{2 N} \in\{-1,1\}
\end{gathered}
$$

## Concluding remarks

- (Sub-)Optimality of diversification under heavy tails \& dependence
- (Non-)robustness of models in economics \& finance to heavy tails, heterogeneity \& dependence
- General representations for joint cdf's and copulas of arbitrary r.v.'s
- Joint cdf's and copulas of dependent r.v.'s $=$ sums of $U$-statistics in independent r.v.'s
- Similar results: expectations of arbitrary statistics in dependent r.v.'s
- New representations for multivariate dependence measures
- Complete characterizations of classes of dependent r.v.'s
- Methods for constructing new copulas
- Modeling different dependence structures


## Concluding remarks

- Copula-based modeling for time series
- Characterizations of dependence in terms of copulas
- Markovness of arbitrary order
- Combining Markovness with other dependencies:
$m$-dependence, $r$-independence, martingaleness, conditional symmetry Non-Markovian processes satisfying Kolmogorov-Chapman SE


## Concluding remarks

- New flexible copulas to combine dependencies
- Expansions by linear functions (Eyraud-Fairlie-Gumbel-Morgensten copulas)
- power functions (power copulas); Fourier polynomials (Fourier copulas)
- Impossibility/reduction: Copula-based dependence + specific copulas $\Leftrightarrow$ Independence


## Copula memory

- Long-memory via copulas: various definitions
- Dependence measures \& copulas
- Gaussian \& EFGM $\Rightarrow$ short-memory Markov
- Fast exponential decay of dependence between $X_{t} \& X_{t+h}$
- Numerical results $\Rightarrow$ Clayton copula-based Markov $\{X t\}$ : can behave as long memory (copulas) in finite samples
- High persistence important for finance \& economics
- Long memory-like: $X_{t} \& X_{t+h}$ : slow decay of dependence for commonly used lages $h$
- Volatility modeling \& Nonlinear dependence in finance
- Non-linear CH \& long memory-like volatility
- Generalizations of GARCH


## Copula memory

Beare (2008) \& Chen, Wu \& Yi (2008): numerical \& theoretical results on (short \& long) memory in copulas

Beare (2008): $\alpha, \beta \& \phi$-mixing

- $\kappa(h) \leq \alpha(h) \leq \beta(h) \leq 0.5 \phi(h)$
- Numerical results $\Rightarrow$ Clayton: exponential decay in $\beta(h) \Rightarrow$ short $\kappa$-memory in copulas

Theoretical results in Chen, Wu \& Yi (2008):

- Clayton: weakly dependent \& short memory in terms of mixing properties!
- Our numerical results + Chen, Wu \& Yi (2008): Non-robustness of procedures for detecting long memory in copulas


## Objectives and key results

- (Sub-)Optimality of diversification under heavy tails \& dependence
- (Non-)robustness of models in economics \& finance to heavy tails, heterogeneity \& dependence
- M. Ibragimov, R. Ibragimov \& J. Walden, Heavy-tailedness and Robustness in Economics and Finance, Lecture Notes in Statistics, Springer, Forthcoming.
- R. Ibragimov \& A. Prokhorov, Topics in Majorization, Stochastic Openings and Dependence Modeling in Economics and Finance, World Scientific Press, In preparation.
- General representations for joint cdf's and copulas of arbitrary r.v.'s
- Copula-based modeling for time series
- Characterizations of time series dependence in terms of copulas
- New flexible copulas to combine dependencies
- Long-memory via copulas: various definitions
- Non-robustness of procedures for detecting long memory in copulas

