Heavy-tailedness and diversification disasters: Implications for models in economics, finance and insurance

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# Objectives and key results

- (Sub-)Optimality of diversification under heavy tails & dependence
- (Non-)robustness of models in economics & finance to heavy tails, heterogeneity & dependence
- Implications for financial & (re-)insurance markets: Diversification traps & disasters
  - M. Ibragimov, R. Ibragimov & J. Walden, *Heavy-tailedness* and *Robustness in Economics and Finance*, Lecture Notes in Statistics, Springer, Forthcoming.
  - R. Ibragimov & A. Prokhorov, Topics in Majorization, Stochastic Openings and Dependence Modeling in Economics and Finance, World Scientific & Imperial College Press, In preparation.

# Imperial College London BUSINESS SCHOOL Stylized Facts of Real-World Returns



# Dependence vs. margins in economic and financial problems

- Problems in finance, economics & risk management: Solution is affected by both
  - Marginal distributions (Heavy-Tailedness, Skewness)
  - Dependence (Positive or Negative, Asymmetry)
- Portfolio choice & value at risk (VaR)
  - Marginal effects under independence: Heavy-Tailedness

Moderately HT vs. extremely HT  $\implies$  Opposite solutions

- Different solutions: Positive vs. negative dependence
- Similar conclusions on (non-)robustness to heavy-tailedness: other models in economics, finance & econometrics:
  - Optimal **bundling**, firm **growth** theory, **efficiency** of statistical & econometric estimators, **time series** models

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Simulated normal and heavy-tailed series



# Heavy-tailed margins

- Many economic & financial time series: power law tails:  $P(|X| > x) \approx \frac{C}{\pi^{\alpha}}, \alpha > 0$ : tail index
- Moments of order  $p \ge \alpha$ : infinite;  $E|X|^p < \infty$  iff  $p < \alpha$ 
  - $\alpha \leq 4 \implies$  Infinite fourth moments:  $EX^4 = \infty$   $\alpha \leq 2 \implies$  Infinite variances:  $EX^2 = \infty$

  - $\alpha \leq 1 \implies$  Infinite first moments:  $E|X| = \infty$
- Returns on many stocks & stock indices:  $\alpha \in (2, 4)$ 
  - $\Rightarrow$  finite variance, infinite fourth moment

# A tale of two tails



Figure: Tails of Cauchy distributions are heavier than those of normal distributions. Tails of Lévy distributions are heavier than those of Cauchy or normal distributions.

# A tale of two tails



Figure: Heavy-tailed distributions: more extreme observations

# Heavy-tailed margins

 $P(|X| > x) \approx \frac{C}{x^{\alpha}}$ 

- Income:  $\alpha \in [1.5, 3] \Rightarrow$  infinite  $EX^4$ , possibly infinite variances
- Wealth:  $\alpha \approx 1.5 \Rightarrow$  infinite variances!
- Returns from technological innovations, Operational risks: α < 1 ⇒ infinite means E|X| = ∞!
- Firm sizes, sizes of largest mutual funds, city sizes:  $\alpha \approx 1$
- Economic losses from earthquakes:  $\alpha \in [0.6, 1.5]$ 
  - $\Rightarrow$  infinite variances, possibly infinite means
- Economic losses from hurricanes:  $\alpha \approx 1.56$ ;  $\alpha \approx 2.49$

### Stable distributions

- $X \sim S_{\alpha}(\sigma)$ : symmetric stable distribution,  $\alpha \in (0, 2]$ CF:  $E(e^{ixX}) = exp\{-\sigma^{\alpha}|x|^{\alpha}\}$ 
  - Normal *N*(0, σ): α = 2
  - Cauchy:  $\alpha = 1$ ,  $f(x) = \frac{\sigma}{\pi(\sigma^2 + x^2)}$

• Lévy: 
$$\alpha = 1/2$$
, support  $[0, \infty)$ ,  $f(x) = \frac{\sigma}{\sqrt{2\pi}} x^{-3/2} \exp(-\frac{1}{2x})$ 

- Power laws:  $P(|X| > x) \approx \frac{c}{x^{\alpha}}, \ \alpha \in (0,2)$ 
  - Moments  $E|X|^p$ : finite iff  $p < \alpha$
  - Infinite variances for  $\alpha < 2$
- **Portfolio** formation:  $\sum_{i=1}^{n} w_i X_i =_d (\sum_{i=1}^{n} w_i^{\alpha})^{1/\alpha} X_1$ •  $\alpha = 2$  (normal):  $\frac{1}{\sqrt{n}} (X_1 + ... + X_n) =_d X_1$

# Value at risk (VaR)

#### VaR

- Risk X; positive values = losses
- Loss probability q
- $VaR_q(X) = z : P(X > z) = q$
- **Risks** *X*<sub>1</sub>, ..., *X<sub>n</sub>*
- $Z_w = \sum_{i=1}^n w_i X_i$ : return on portfolio with weights  $w = (w_1, ..., w_n)$
- Problem of interest:

Minimize  $VaR_q(Z_w)$ 

- s.t.  $w_i \ge 0$ ,  $\sum_{i=1}^n w_i = 1$
- When diversification ⇒ decrease in portfolio riskiness (VaR)?

# **Diversification & risk**

- Most diversified:  $\underline{w} = (1/n, 1/n, ..., 1/n) \Rightarrow Z_{\underline{w}} = \frac{1}{n} \sum_{i=1}^{n} X_i$
- Least diversified:  $\overline{w} = (1, 0, ..., 0) \Rightarrow Z_{\overline{w}} = X_1$

• 
$$Z_{\underline{w}} = \frac{1}{n} \sum_{i=1}^{n} X_i =_d \frac{1}{\sqrt{n}} X_1 = \frac{1}{\sqrt{n}} Z_{\overline{w}}$$

- $VaR_q(Z_{\underline{w}}) = \frac{1}{\sqrt{n}} VaR_q(Z_{\overline{w}}) < VaR_q(Z_{\overline{w}})$
- VaR<sub>q</sub>(Z<sub>w</sub>) : ∖ as n ∧ (Diversification ∧)

# **Diversification & risk**

- $X_1,...,X_n\sim S_{1/2}(\sigma)$ , lpha=1/2, Lévy distribution
  - $Z_{\underline{w}} = \frac{1}{n} \sum_{i=1}^{n} X_i =_d \left[ \sum_{i=1}^{n} (\frac{1}{n})^{1/2} \right]^2 X_1 = n X_1 = n Z_{\overline{w}}$
  - $VaR_q(Z_{\underline{w}}) = nVaR_q(Z_{\overline{w}}) > VaR_q(Z_{\overline{w}})$
  - $VaR_q(Z_{\underline{w}})$  :  $\nearrow$  as  $n \nearrow$  (Diversification  $\nearrow$ )
- Heavy tails (margins) matter:

diversification  $\Longrightarrow$  opposite effects on portfolio riskiness

• Skewness: typically priced

# Heavy-tailedness & diversification

• Moderate heavy tails  $\alpha > 1$ : finite first moments

 $VaR_q(Z_{w}) < VaR_q(Z_{\overline{w}}) \ \forall q > 0$ 

Optimal to diversify for all loss probabilities q

• Extremely heavy tails  $\alpha < 1$  : infinite first moments

 $VaR_q(Z_{w}) < VaR_q(Z_{\overline{w}}) \ \forall q > 0$ 

Diversification: suboptimal for all loss probabilities q

• Similar conclusions: Many other models in economics & finance

- Firm growth theory, optimal bundling, monotone consistency of sample mean, efficiency of linear estimators
- Robust to moderate heavy tails
- Properties: reversed under extremely heavy tails

# What happens for intermediate heavy-tails?

•  $X_1, ..., X_n$  i.i.d. stable with  $\alpha = 1$ : Cauchy distribution

• Density 
$$f(x) = \frac{\sigma}{\pi(\sigma^2 + x^2)}$$

- Heavy power law tails:  $P(|X| > x) \approx \frac{C}{x}$
- Infinite first moment

• 
$$Z_w = \sum_{i=1}^n w_i X_i =_d X_1 \ \forall w = (w_1, ..., w_n) : w_i \ge 0,$$

• Diversification: no effect at all!

# Summary so far: Diversification for heavy-tailed and bounded distributions



Figure: N = 10 risks/insurer; M = 7 insurers

• D: Individual/non-diversification corners vs insurer and reinsurer equilibrium

### 1st example: full risk pooling with normally distributed risks



### 2nd example: Bernoulli-Lévy distribution with limited liability



# Implications for markets for catastrophic risks

- Equilibria in re-insurance markets for catastrophe risks (Ibragimov, Jaffee and Walden, RFS)
  - A diversification equilibrium with full risk pooling for normally distributed (light-tailed) risks
  - No risk pooling & no insurance or reinsurance activity (market collapse) for extremely heavy-tailed cat risks
  - Intermediate cases (heavy tails): both
    - Diversification equilibria, in which insurers offer catastrophe coverage and reinsure their risks
    - Non-diversification equilibria with no insurance or re-insurance
    - A coordination problem must be solved to shift from the bad to the good equilibrium

Government regulations or well functioning capital markets

# Implications for markets for catastrophic risks

- Catastrophic risks have many features favorable to the provision of insurance
  - Generally independent over risk types and geography
  - Few issues of asymmetric information at the risk level
  - So a complete failure of these markets is puzzling
- We have shown that market failures (non-diversification traps) may arise when risks are fat-tailed and there is limited liability
  - Diversification may not be beneficial for the single insurer, although a full reinsurance equilibrium may exist.
  - Government programs (or diversified equity owners) may allow the system to reach the full diversification outcome

# **Diversification & dependence**

- Minimize  $VaR_q(w_1X_1 + w_2X_2)$  s.t.  $w_1, w_2 \ge 0, w_1 + w_2 = 1$
- Independence:
  - Optimal portfolio: (*w˜*<sub>1</sub>, *w˜*<sub>2</sub>) = (<sup>1</sup>/<sub>2</sub>, <sup>1</sup>/<sub>2</sub>) (diversified) if α > 1 (not extremely heavy-tailed, finite means)
  - (*w˜*<sub>1</sub>, *w˜*<sub>2</sub>) = (1,0) (not diversified, one risk) if α < 1 (extremely heavy-tailed, infinite means)</li>

# **Diversification & dependence**

- Extreme positive dependence:  $X_1 = X_2$  (a.s.) comonotonic risks
  - $VaR_q(w_1X_1 + w_2X_2) = VaR_q(X_1) \ \forall w$
  - Diversification: no effect at all (similar to Cauchy) regardless of heavy-tailedness
- Extreme negative dependence  $X_1 = -X_2$  (a.s.) countermonotonic risks
  - $VaR_q(w_1X_1 + w_2X_2) = (w_1 w_2)VaR_q(X_1)$
  - Optimal portfolio:  $\underline{w} = (1/2, 1/2)$  (most diversified regardless of heavy-tailedness
- Optimal portfolio choice: affected by both dependence & properties of margins

# **Copulas and dependence**

- Main idea: separate effects of dependence from effects of margins
  - What matters more in portfolio choice: heavy-tailedness & skewness or (positive or negative) dependence?
- Copulas: functions that join together marginal cdf's to form multidimensional cdf

# **Copulas and dependence**

- Sklar's theorem
- Risks X, Y:
  - Joint cdf  $H_{XY}(x, y) = P(X \le x, Y \le y)$ : affected by dependence and by marginal cdf's  $F_X(x) = P(X \le x)$  and  $G_Y(x) = P(Y \le y)$

$$H_{XY}(x,y) = \underbrace{C_{XY}}_{} (\underbrace{F_X(x), G_Y(y)}_{})$$

dependence marginals

• C<sub>XY</sub>: captures all dependence between risks X and Y

# **Copulas and dependence**

#### Advantages:

- Exists for any risks (correlation: finiteness of second moments)
- Characterizes all dependence properties
- Flexibility in dependence modeling
  - Asymmetric dependence: Crashes vs. booms
  - Positive vs. negative dependence
  - Independence: Nested as a particular case: Product copula, particular values of parameter(s)
  - Extreme dependence: X = Y or X = −Y ⇔ extreme copulas; dependence in C<sub>XY</sub> varies in between

# Copula structures

• Archimedean copulas

$$C(u, v) = \phi^{-1}(\phi(u) + \phi(v))$$

### • Contagion: Non-zero tail dependence coeff.

$$\lambda_{L} = \lim_{u \to 0+} P[Y \le F^{-1}(u) | X \le F_{X}^{-1}(u)] = \lim_{u \to 0+} \frac{C(u, u)}{u}$$
$$\lambda_{U} = \lim_{u \to 1-} P[Y > F^{-1}(u) | X > F_{X}^{-1}(u)] = \lim_{u \to 1-} \frac{1 - 2u + C(u, u)}{1 - u}$$

### • Clayton & Gumbel copulas

# Copula structures

• Eyraud-Farlie-Gumbel-Morgenstern (EFGM):

$$C(u,v) = uv[1+\gamma(1-u)(1-v)]$$

 $\gamma \in [-1,1]$  : dependence parameter Tail independent: no contagion

• Heavy-tailed Pareto marginals:

$$egin{aligned} P(X > x) &= rac{1}{x^lpha}, \ x \geq 1 \ P(X > x) &= rac{1}{x^lpha}, \ x \geq 1 \end{aligned}$$

• Power laws, tail index  $\alpha$ 

# Diversification: Copulas & heavy tails

Embrechts, Nešlehová & Wüthrich (2009): Archimedean copulas

• Moderate heavy tails  $\alpha > 1$  : finite first moment

$$VaR_q(\frac{X+Y}{2}) < VaR_q(X)$$
 for sufficiently small q

**Optimal** to **diversify** for **sufficiently small** loss probabilities q

• Extremely heavy tails  $\alpha < 1$ : infinite first moments

$$VaR_q(rac{X+Y}{2}) > VaR_q(X)$$
 for sufficiently small  $q$ 

Diversification: suboptimal for suff. small loss prob. q

Ibragimov & Prokhorov (2013): Similar conclusions for EFGM

• Tail independent EFGM & tail dependent Archimedean (Clayton, Gumbel): same boundary  $\alpha = 1$  as in the case of independence

### When dependence helps: Student-t copulas

• Conclusions similar to independence: Models with common shocks

$$X_1 = ZY_1, X_2 = ZY_2, ..., X_n = ZY_n$$

- Common shock Z > 0 affecting all risks X<sub>1</sub>,..., X<sub>n</sub>
- $Y_1, ..., Y_n$ : i.i.d. normal or heavy-tailed with tail index  $\alpha$

*Z* : heavy-tailed with tail index  $\beta$ 

Then  $X_i$ : heavy-tailed with tail index  $\gamma = \min(\alpha, \beta)$ 

- Important particular case: (Dependent) Multivariate Student-t $X_1, X_2, ..., X_n$  with  $\alpha$  d.f. (tail index)  $\Rightarrow$  Optimal to diversify for all loss probabilities q regardless of tail index  $\alpha$ 
  - Tail dependent Student-t copula and heavy-tailed margins with arbitrary tail index  $\alpha$  : diversification pays off
- Contrast: Independent Student-t X<sub>1</sub>, X<sub>2</sub>, ..., X<sub>n</sub> with α d.f. (tail index): diversification optimal for α > 1; suboptimal for α < 1</li>

# Diversification: Heavy-tailedness & dependence matter

- **Independence**, **Tail dependent** models with **common shocks** (e.g., Student-*t* distr. = Student-*t* copula with Student-*t* marginals):
  - Diversification always pays off for all loss probabilities q
- Tail independent EFGM, possibly tail dependent Archimedean copulas (e.g., Clayton & Gumbel):
  - Dividing boundary  $\alpha = 1$  for sufficiently small loss probability q
- Numerical results on interplay of heavy-tailedness & dependence (copula) assumptions and loss probability q in diversification decisions:
  - Deviations from threshold  $\alpha=1$  for different copulas and loss probabilities q
- Theoretical results for general copulas = ?
- (Non-)robustness of other models in economics & finance

# Key results

- (Sub-)Optimality of diversification under heavy tails & dependence
- (Non-)robustness of models in economics & finance to heavy tails, heterogeneity & dependence
- Implications for financial & (re-)insurance markets: Diversification traps & disasters
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# Characterizations of copulas & dependence

- $V_1, ..., V_n$ : i.i.d.  $\mathcal{U}([0, 1])$
- C: n-copula iff  $\exists \tilde{g}_{i_1,...,i_c}$  s.t.
  - A1 (integrability):

$$\int_{0}^{1}...\int_{0}^{1}| ilde{g}_{i_{1},...,i_{c}}(t_{i_{1}},...,t_{i_{c}})|dt_{i_{1}}...dt_{i_{c}}<\infty$$

A2 (degeneracy):

$$E_{V_{i_k}}\left[\tilde{g}_{i_1,...,i_c}(V_{i_1},...,V_{i_{k-1}},V_{i_k},V_{i_{k+1}},...,V_{i_c})\right] = 0$$

A3 (positive definiteness):

$$ilde{U}_n(V_1,...,V_n) \equiv \sum_{c=2}^n \sum_{1 \le i_1 < ... < i_c \le n} ilde{g}_{i_1,...,i_c}(V_{i_1},...,V_{i_c}) \ge -1$$

• **Representation for** *C* :

$$C(u_1,...,u_n) = \int_0^{u_1} ... \int_0^{u_n} (1 + \tilde{U}_n(t_1,...,t_n)) \prod_{i=1}^n dt_i$$

• 
$$\tilde{U}_n$$
: sum of degenerate U-statistics

Device for constructing *n*-copulas and cdf's

#### • Bivariate Eyraud-Farlie-Gumbel-Morgenstern copulas & cdf's:

$$C_{\theta}(u,v) = uv \left(1 + \theta(1-u)(1-v)\right)$$
$$H_{\theta}(x,y) = F(x)G(y)\left(1 + \theta(1-F(x))(1-G(y))\right)$$

$$n = 2; \ \tilde{g}_{1,2}(t_1, t_2) = \theta(1 - 2t_1)(1 - 2t_2), \ \theta \in [-1, 1]$$

• Multivariate EFGM copulas & cdf's:

$$C_{\theta}(u_1, u_2, ..., u_n) = \prod_{i=1}^n u_i \left( 1 + \theta \prod_{i=1}^n (1 - u_i) \right)$$
$$\tilde{g}_{i_1, ..., i_c}(t_{i_1}, ..., t_{i_c}) = \theta_{i_1, ..., i_c} (1 - 2t_{i_1}) (1 - 2t_{i_2}) ... (1 - 2t_{i_c})$$

• Generalized multivariate EFGM copulas (Johnson and Kotz, 1975, Cambanis, 1977)

$$C(u_1, ..., u_n) = \prod_{k=1}^n u_k \left( 1 + \sum_{c=2}^n \sum_{1 \le i_1 < ... < i_c \le n} \theta_{i_1, ..., i_c} (1 - u_{i_k}) \right)$$
$$\tilde{g}_{i_1, ..., i_c}(t_{i_1}, ..., t_{i_c}) = 0, \ c < n - 1$$
$$\tilde{g}_{1, 2, ..., n}(t_1, t_2, ..., t_n) = \theta(1 - 2t_1)(1 - 2t_2)...(1 - 2t_n)$$

 Generalized EFGM copulas: complete characterization of joint cdf's of two-valued r.v.'s (Sharakhmetov & Ibragimov, 2002)

# From dependence to independence through *U*-statistics

 $\mathcal{G}_n$ : sums of *U*-statistics

$$U_n(\xi_1,...,\xi_n) = \sum_{c=2}^n \sum_{1 \le i_1 < ... < i_c \le n} g_{i_1,...,i_c}(\xi_{i_1},...,\xi_{i_c})$$

 $g_{i_1,...,i_c}$ : satisfy A1-A3

• Arbitrarily dependent r.v.'s:

sum of U-statistics in independent r.v.'s

with canonical kernels

- Reduction of problems for dependence to well-studied objects
- Transfer of results for U-**statistics** under

#### independence

# From dependence to independence through *U*-statistics

- X<sub>1</sub>,..., X<sub>n</sub>: **1-cdf's** F<sub>k</sub>(x<sub>k</sub>)
- $\xi_1, ..., \xi_n$ : independent copies (1-cdf's  $F_k(x_k)$ )

 $\exists U_n \in \mathcal{G}_n \text{ s.t. } \forall f : \mathbf{R}^n \to \mathbf{R}$ 

$$Ef(X_1,...,X_n) = Ef(\xi_1,...,\xi_n) \Big( 1 + U_n(\xi_1,...,\xi_n) \Big)$$

• Representation for c.f.'s:

$$Eexp\left(i\sum_{k=1}^{n}t_{k}X_{k}\right) = Eexp\left(i\sum_{k=1}^{n}t_{k}\xi_{k}\right) + \\Eexp\left(i\sum_{k=1}^{n}t_{k}\xi_{k}\right)U_{n}(\xi_{1},...,\xi_{n})$$

‡ CLT for bivariate r.v.'s

# **Characterizations of dependence**

# • Canonical g's: complete characterizations of

dependence properties

•  $X_1, ..., X_n$ : r-independent if  $\forall r$  jointly independent  $\Leftrightarrow$  $g_{i_1,...,i_c}(V_{i_1}, ..., V_{i_c}) = 0$  (a.s.)  $1 \le i_1 < ... < i_c \le n, c = 2, ..., r$ 

$$g_{i_1,...,i_{r+1}}(u_{i_1},...,u_{i_{r+1}}) = \frac{\alpha_1...\alpha_n}{\alpha_{i_1}...\alpha_{i_{r+1}}} \left( (k+1)u_{i_1}^k - (k+2)u_{i_1}^{k+1} \right) \times ... \times \left( (k+1)u_{i_c}^k - (k+2)u_{i_c}^{k+1} \right)$$

$$C(u_1, ..., u_n) = \prod_{i=1}^n u_i \left( 1 + \sum_{1 \le i_1 < \dots < i_{r+1} \le n} \frac{\alpha_1 \dots \alpha_n}{\alpha_{i_1} \dots \alpha_{i_{r+1}}} \times (u_{i_1}^k - u_{i_1}^{k+1}) \times \dots \times (u_{i_{r+1}}^k - u_{i_{r+1}}^{k+1}) \right)$$

Extensions of Wang (1990) (k = 0)

# **Copulas and Markov processes**

• Darsow, Nguyen and Olsen, 1992: copulas and first-order Markovness

•  $A, B : [0,1]^2 \rightarrow [0,1]$  :

$$(A * B)(x, y) = \int_0^1 \frac{\partial A(x, t)}{\partial t} \cdot \frac{\partial B(t, y)}{\partial t} dt$$

•  $A: [0,1]^m \rightarrow [0,1], \ B: [0,1]^n \rightarrow [0,1]: \star - \text{product}$ 

$$A \star B(x_1, \dots, x_{m+n-1}) =$$

$$\int_0^{x_m} \frac{\partial A(x_1, \dots, x_{m-1}, \xi)}{\partial \xi} \cdot \frac{\partial B(\xi, x_{m+1}, \dots, x_{m+n-1})}{\partial \xi} d\xi$$

# **Copulas and Markov processes**

#### • Transition probabilities

$$P(s, x, t, A) = P(X_t \in A | X_s = x)$$
 satisfy CKE's

 $\text{iff } C_{st} = C_{su} \ast C_{ut} \ \forall s < u < t$ 

• X<sub>t</sub>: first-order Markov iff

$$C_{t_1,\ldots,t_n}=C_{t_1t_2}\star C_{t_2t_3}\star\ldots\star C_{t_{n-1}t_n}$$

# New results: Higher-order Markovness and copulas

•  $\{X_t\}_{t\in\mathcal{T}}$ : *k*-order Markov  $\Leftrightarrow$ 

$$egin{aligned} & Pig(X_t < x_t ig| X_{t_1}, ..., X_{t_{n-k}}, X_{t_{n-k+1}}, ..., X_{t_n}ig) = \ & Pig(X_t < x_t ig| X_{t_{n-k+1}}, ..., X_{t_n}ig) \end{aligned}$$

• Complete characterization in terms of (k + 1)-copulas

•  $C_{t_1,...,t_k}$ : copulas of  $X_{t_1},...,X_{t_k}$ 

•  $\{X_t\}_{t \in T}$ : k-order Markov iff  $\forall t_1 < ... < t_n, n \ge k+1$ 

$$C_{t_1,...,t_n} = C_{t_1,...,t_{k+1}} \star^k C_{t_2,...,t_{k+2}} \star^k ... \star^k C_{t_{n-k},...,t_n}$$

# **Stationary case**

#### • $X_t$ : stationary k-order Markov iff

$$C_{1,...,n}(u_1,...,u_n) = C \star^k C \star^k ... \star^k C(u_1,...,u_n)$$
  
=  $C^{n-k+1}(u_1,...,u_n) \quad \forall n \ge k+1$ 

C: (k+1)- copula s.t.

$$C_{i_1+h,...,i_l+h} = C_{i_1,...,i_l}, \ 1 \le j_1 < ... < j_l \le k+1$$

•  $C^s$ : *s*-fold product  $\star^k$  of *C* 

# Advantages of copula-based approach

• Modeling higher order Markov processes alternative to transition matrices

‡ Instead of initial distribution & transition probabilities:

Prescribe marginals & (k + 1)-copulas

Generate copulas of higher order & finite-dimensional cdf's

<sup>‡</sup> Advantage: separation of properties of marginals (fat-tailedness) & dependence properties (conditional symmetry, m-dependence, r-independence, mixing)

# Advantages of copula-based approach

• Inversion method:

New k-Markov with dependence similar to a given Markov process Different marginals

 $\ddagger X_t: \text{ stationary } k-\text{Markov}$  $(k+1)-\text{cdf } \tilde{F}(x_1,...,x_{k+1}), \ 1-\text{cdf } F$ 

 $\Rightarrow$  (k + 1)-copula:

$$C(u_1,...,u_{k+1}) = \tilde{F}\Big(F^{-1}(u_1),...,F^{-1}(u_{k+1})\Big)$$

<sup>†</sup> Another 1–cdf *G*: **Stationary** k–**Markov**, **same** dependence as { $X_t$ }, **different** 1-marginal *G*:

(k+1)-copula:

$$C(u_1,...,u_{k+1}) = \tilde{F}\Big(G^{-1}(u_1),...,G^{-1}(u_{k+1})\Big)$$

Representation  $\Rightarrow$  Higher-order copulas & cdf's

 $\{X_t\}$ : stationary *C*-based *k*-Markov chain

# Advantages of copula-based approach

#### • C: all dependence properties of the time series

 $\ddagger k$ -independence, m-dependence, martingaleness, symmetry

<sup>‡</sup> On-going project with Johan Walden: characterizations of **time-irreversibility**; focus on  $C_{t_1,...,t_k} = C_{t_k,...,t_1}$ 

‡ Applications: forward-looking vs. backward-looking market participants ("fundamentalists" vs. noise traders or "chartists")

‡ "Compass rose" for  $P_{t-1}$  and  $P_t$ : symmetry in copulas

# Combining higher-order Markovness with other dependence properties

• A number of studies in **dependence modeling: Higher-order Markovness** + *m*-**dependence** & *r*-**independence** 

Lévy (1949): 2nd order Markovness + pairwise independence

Rosenblatt & Slepian (1962): N-order N-independent stationary Markov

#### • Impossibility/reduction :

*N*-order Markov + *N*-independence + two-valued  $\Leftrightarrow$  joint independence **‡** Testing sensitivity to WD in DGP Rosenblatt & Slepian (1962)

# Combining Markovness with other dependencies

### ‡ Examples:

Not 1-order Markovian

But 1-st order transition probabilities

 $P(s, x, t, A) = P(X_t \in A | X_s = x)$  satisfy C-K SE

$$P(s,x,t,A) = \int_{-\infty}^{\infty} P(u,\xi,t,A)P(s,x,u,d\xi)$$

(other examples: Feller, 1959, Rosenblatt, 1960)

# Combining Markovness with other dependencies

‡ 1-dependent Markov: Aaronson, Gilat and Keane (1992)

Burton, Goulet and Meester (1993), Matúš (1996)

‡ Matúš (1998): m-dependent discrete-space Markov

# ‡ Impossibility/Reduction:

 $\nexists$  stationary *m*-dependent Markov if *card*( $\Omega$ ) < *m* + 2

# Markovness of higher-order and *k*-independence

- Characterization of stationary
- k-independent k-Markov processes
- $\{X_t\}$ : *C*-based *k*-independent stationary
- k-Markov iff

$$\frac{\partial^{k+1}C(u_1,...,u_{k+1})}{\partial u_1...\partial u_{k+1}} = 1 + g(u_1,...,u_{k+1})$$

 $g: [0,1]^{k+1} \rightarrow [0,1]$ : canonical g-function

(Integrability + more degeneracy + positive definiteness)

### Markovness of higher-order and *k*-independence

$$\begin{split} \int_0^1 ... \int_0^1 |g(u_1,...,u_{k+1})| du_1 ... du_{k+1} < \infty \\ \int_0^1 ... \int_0^1 g(u_1,...,u_{k+1}) g(u_2,...,u_{k+2}) ... g(u_s,...,u_{k+s}) du_{i_1} ... du_{i_s} = 0 \\ \forall s \le u_{i_1} < ... < u_{i_s} \le k+1, \ s = 1,2,..., \left[\frac{k+1}{2}\right] \\ g(u_1,...,u_{k+1}) \ge -1 \end{split}$$

• Integration: w.r. to all s among  $u_s, u_{s+1}, ..., u_{k+1}$  common to all g-functions  $g(u_1, ..., u_{k+1}), g(u_2, ..., u_{k+2}), ..., g(u_s, ..., u_{k+s})$ 

- *k*-marginals: product copulas, independence
- k-independence: satisfied

### Markovness of higher-order and *m*-independence

•  $\{X_t\}$ : C-based m-dependent 1-Markov iff

$$\frac{\partial^2 C(u_1, u_2)}{\partial u_1 \partial u_2} = 1 + g(u_1, u_2)$$

 $g: [0,1]^2 \rightarrow [0,1]$ : canonical g-function:

$$\int_0^1 \int_0^1 |g(u_1, u_2)| du_1 du_2 < \infty$$
  
 $\int_0^1 g(u_1, u_2) du_i = 0, \ g(u_1, u_2) \ge -1$   
 $\int_0^1 g(u_1, u_2) g(u_2, u_3) ... g(u_m, u_{m+1}) du_2 du_3 ... du_m = 0$ 

‡ Integration: w.r. to  $u_2, u_3, \dots, u_m$  more than once among  $g(u_1, u_2), g(u_2, u_3), \dots, g(u_m, u_{m+1})$ 

 $X_1$ ,  $X_{m+1}$ : independent; Process: *m*-dependent

# New examples via existing constructions

- Higher-order Markovness + martingaleness
- Inversion method + existing examples  $\Rightarrow$
- k-independent, m-dependent Markov processes

different marginals

# Reduction & impossibility for *k*-order Markov processes

$$\ddagger \frac{\partial^{k+1} C(u_1,...,u_{k+1})}{\partial u_1...\partial u_{k+1}} = 1 + g(u_1,...,u_{k+1})$$

 $\ddagger g :$  product form (EFGM-type):  $g(u_1, u_2, ..., u_{k+1}) = \alpha f(u_1) f(u_2) ... f(u_{k+1})$ 

 $\Leftrightarrow \{X_t\}$ : jointly independent

# Examples: EFGM and power copulas

• (k+1)-**EFGM** copulas:

$$C(u_1, u_2, ..., u_{k+1}) = \prod_{i=1}^{k+1} u_i \Big( 1 + \alpha (1 - u_1)(1 - u_2) ... (1 - u_{k+1}) \Big)$$

$$g(u_1, u_2, ..., u_{k+1}) = \alpha(1 - 2u_1)(1 - 2u_2)...(1 - 2u_{k+1})$$

• (k+1)-power copulas

$$C(u_1, u_2, ..., u_{k+1}) = \prod_{i=1}^{k+1} u_i \Big( 1 + \alpha (u_1' - u_1'^{i+1}) (u_2' - u_2'^{i+1}) ... (u_{k+1}' - u_{k+1}'^{i+1}) \Big)$$

 $l \ge 0$  (EFGM: l = 0)

# Impossibility/reduction for *m*-dependence

• 
$${X_t}: C-based m-dependent Markov$$

$$\ddagger \frac{\partial^2 C(u_1, u_2)}{\partial u_1 \partial u_2} = 1 + \alpha f(u_1) f(u_2)$$

(separable product form)

- $\Leftrightarrow X_t$ : jointly independent
- Representations  $\Rightarrow$

$$\int_{0}^{1} \dots \int_{0}^{1} \alpha^{m} f(u_{1}) f^{2}(u_{2}) \dots f^{2}(u_{m}) f(u_{m+1}) du_{2} \dots du_{m} = 0;$$
  
$$\alpha^{m} f(u_{1}) f(u_{m+1}) \Big[ \int_{0}^{1} f^{2}(u_{2}) du_{2} \Big]^{m-1} = 0$$

 $\Rightarrow f = 0 \Leftrightarrow$  Independence

# Examples, new and old

**‡ EFGM copulas**, k = 1:

$$C(u_1, u_2) = u_1 u_2 \Big( 1 + \alpha (1 - u_1) (1 - u_2) \Big)$$
$$g(u_1, u_2) = \alpha (1 - 2u_1) (1 - 2u_2)$$

• Limitations of EFGM copulas,

separable copulas:

Complement & generalize existing results

# Examples, new and old

# ‡ Cambanis (1991): common dependenciescannot be exhibited by multivariate EFGM

$$egin{aligned} & C_{j_1,...,j_n}(u_{j_1},...,u_{j_n}) = \ & \prod_{s=1}^n u_{j_k} \Big( 1 + \sum_{1 \leq l < m \leq n} lpha_{lm} (1-u_{j_l}) (1-u_{j_m}) \Big) \end{aligned}$$

‡ Rosenblatt & Slepian (1962): non-existence of bivariate N-independent N-Markov

Sharakhmetov & Ibragimov (2002):

EFGM copulas for two-valued r.v.'s

**† Technical difficulties** in modeling

# Solution: New flexible copula classes

- Copula-based TS with flexible dependencies
- ‡ Copulas based on Fourier polynomials
- k-independent k-Markov: Conditions satisfied for

$$g(u_1, ..., u_{k+1}) = \sum_{j=1}^{N} \left[ \alpha_j \sin(2\pi \sum_{i=1}^{k+1} \beta_i^j u_i) + \gamma_j \cos(2\pi \sum_{i=1}^{k+1} \beta_i^j u_i) \right]$$
  

$$\ddagger \alpha_j, \gamma_j \in \mathbf{R}, \ \beta_i^j \in \mathbf{Z}, \ i = 1, ..., k+1, \ j = 1, ..., N:$$
  

$$\ddagger \beta_1^{j_1} + \sum_{l=2}^{s} \epsilon_{l-1} \beta_l^{j_l} \neq 0$$
  

$$\epsilon_1, ..., \epsilon_{s-1} \in \{-1, 1\}, \ s = 2, ..., k+1$$
  

$$\ddagger 1 + \sum_{j=1}^{N} \left[ \alpha_j \epsilon_j + \gamma_j \epsilon_{j+N} \right] \ge 0, \ \epsilon_1, ..., \epsilon_{2N} \in \{-1, 1\}$$

# **Fourier copulas**

$$C(u_1,...,u_{k+1}) = \int_0^{u_1} \dots \int_0^{u_{k+1}} (1 + g(u_1,...,u_{k+1})) du_1 \dots du_{k+1}$$

(k+1)-Fourier copulas

# **Fourier copulas**

#### • 1-dependent 1-Markov:

Conditions satisfied for Fourier copulas

$$C(u_1, u_2) = \int_0^{u_1} \int_0^{u_2} (1 + g(u_1, u_2)) du_1 du_2$$

$$g(u_1, u_2) = \sum_{j=1}^{N} \left[ \alpha_j sin(2\pi(\beta_1^j u_1 + \beta_2^j u_2)) + \gamma_j cos(2\pi(\beta_1^j u_1 + \beta_2^j u_2)) \right]$$

 $\ddagger \alpha_j, \gamma_j \in \mathbf{R}, \ \beta_1^j, \beta_2^j \in \mathbf{Z}$ :

$$\begin{aligned} \beta_1^{j_1} + \beta_2^{j_2} \neq \mathbf{0} \\ \beta_1^{j_1} - \beta_2^{j_2} \neq \mathbf{0} \\ \mathbf{1} + \sum_{j=1}^N [\alpha_j \epsilon_j + \gamma_j \epsilon_{j+N}] \geq \mathbf{0} \end{aligned}$$

 $\forall \epsilon_1,...,\epsilon_{2\textit{N}} \in \{-1,1\}$ 

# **Concluding remarks**

- (Sub-)Optimality of diversification under heavy tails & dependence
- (Non-)robustness of models in economics & finance to heavy tails, heterogeneity & dependence
- · General representations for joint cdf's and copulas of arbitrary r.v.'s
  - Joint cdf's and copulas of **dependent** r.v.'s = sums of *U*-statistics in **independent** r.v.'s
  - Similar results: expectations of arbitrary statistics in dependent r.v.'s
  - New representations for multivariate dependence measures
  - Complete characterizations of classes of dependent r.v.'s
  - Methods for constructing new copulas
  - Modeling different dependence structures

# **Concluding remarks**

- Copula-based modeling for time series
- Characterizations of dependence in terms of copulas
  - Markovness of arbitrary order
  - Combining Markovness with other dependencies:

m-dependence, r-independence, martingaleness, conditional symmetry

Non-Markovian processes satisfying Kolmogorov-Chapman SE

# **Concluding remarks**

- New flexible copulas to combine dependencies
- Expansions by linear functions (Eyraud-Fairlie-Gumbel-Morgensten copulas)
- power functions (power copulas); Fourier polynomials (Fourier copulas)
- Impossibility/reduction: Copula-based dependence + specific copulas
   ⇔ Independence

# **Copula memory**

- Long-memory via copulas: various definitions
- Dependence measures & copulas
- Gaussian & EFGM ⇒ short-memory Markov
- Fast exponential decay of dependence between X<sub>t</sub> & X<sub>t+h</sub>
- Numerical results ⇒ Clayton copula-based Markov {*Xt*} : can behave as long memory (copulas) in finite samples
  - High persistence important for finance & economics
- Long memory-like: X<sub>t</sub> & X<sub>t+h</sub> : slow decay of dependence for commonly used lages h
- Volatility modeling & Nonlinear dependence in finance
- Non-linear CH & long memory-like volatility
- Generalizations of GARCH

# **Copula memory**

Beare (2008) & Chen, Wu & Yi (2008): numerical & theoretical results on (short & long) memory in copulas

Beare (2008):  $\alpha$ ,  $\beta$  &  $\phi$ -mixing

- $\kappa(h) \leq \alpha(h) \leq \beta(h) \leq 0.5\phi(h)$
- Numerical results  $\Rightarrow$  Clayton: exponential decay in  $\beta(h) \Rightarrow$  short  $\kappa$ -memory in copulas

Theoretical results in Chen, Wu & Yi (2008):

- Clayton: weakly dependent & short memory in terms of mixing properties!
- Our numerical results + Chen, Wu & Yi (2008): Non-robustness of procedures for detecting long memory in copulas

# **Objectives and key results**

- (Sub-)Optimality of diversification under heavy tails & dependence
- (Non-)robustness of models in economics & finance to heavy tails, heterogeneity & dependence
  - M. Ibragimov, R. Ibragimov & J. Walden, *Heavy-tailedness and Robustness in Economics and Finance*, Lecture Notes in Statistics, Springer, Forthcoming.
  - R. Ibragimov & A. Prokhorov, *Topics in Majorization, Stochastic Openings and Dependence Modeling in Economics and Finance*, World Scientific Press, In preparation.
- General representations for joint cdf's and copulas of arbitrary r.v.'s
- Copula-based modeling for time series
- Characterizations of time series dependence in terms of copulas
- New flexible copulas to combine dependencies
- Long-memory via copulas: various definitions
- · Non-robustness of procedures for detecting long memory in copulas